

Kidney Mass and Relative Medullary Thickness of Rodents in Relation to Habitat, Body Size, and Phylogeny

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ABSTRACT

We tested the hypotheses that relative medullary thickness (RMT) and kidney mass are positively related to habitat aridity in rodents, after controlling for correlations with body mass. Body mass, mass-corrected kidney mass, mass-corrected RMT, mass-corrected maximum urine concentration, and habitat (scored on a semiquantitative scale of 1–4 to indicate increasing aridity) all showed statistically significant phylogenetic signal. Body mass varied significantly among habitats, with the main difference being that aquatic species are larger than those from other habitats. Mass-corrected RMT and urine concentration showed a significant positive correlation ($N = 38$; conventional $r = 0.649$, phylogenetically independent contrasts [IC] $r = 0.685$), thus validating RMT as a comparative index of urine concentrating ability. RMT scaled with body mass to an exponent significantly less than 0 ($N = 141$ species; conventional allometric slope = -0.145 [95% confidence interval (CI) = $-0.172, -0.117$], IC allometric slope = -0.132 [95% CI = $-0.180, -0.083$]). Kidney mass scaled to an exponent significantly less than unity ($N = 104$ species; conventional slope = 0.809 [95% CI = $0.751, 0.868$], IC slope = 0.773 [95% CI = $0.676, 0.871$]). Both conventional and phylogenetic analysis indicated that RMT varied among habitats, with ro-

dent from arid areas having the largest values of RMT. A phylogenetic analysis indicated that mass-corrected kidney mass was positively related to habitat aridity.

Introduction

Mammalian kidneys have a dominant role in controlling both the volume and concentration of body fluids. The nephron is the functional unit of the kidney and consists of a glomerulus and well-developed loops of Henle. The morphological and vascular organizations of nephrons enable mammals to produce urine that is significantly more concentrated than their own plasma. Some of these nephrons, “long looped nephrons,” are characterized by an extended renal medullary papilla (Folk 1974; Bankir and de Rouffignac 1985) that reflects the great length of the loop of Henle. The maximum length of the loop of Henle is directly proportional to medullary thickness (Beuchat 1990, 1993, 1996). Sperber’s (1944) work on mammalian kidneys showed a relationship between length of the renal papilla and the availability of drinking water in the natural habitat. Specifically, mammals from arid and semiarid habitats tended to have exceptionally long loops of Henle, as compared with mammals from mesic habitats. Sperber (1944) also proposed the relative medullary thickness as a structural index for quantifying the relative length of the longest loops of Henle. Relative medullary thickness (RMT) is calculated as $(MT/KS) \times 10$, where MT (typically in mm) is the total thickness of the medulla and KS is kidney size (in mm), computed as the cube root of the product of the three linear dimensions of the kidney. Sperber (1944) found that mammals living in arid areas had higher values of RMT than those of similar-sized mammals from more mesic habitats.

According to Gottschalk (1987), Sperber was by no means the first to relate the length of the loop of Henle to urine concentrating ability in mammals. Peter (1909, cited in Gottschalk 1987) had pointed out such a correlation among various species of mammals. The relationship was further evaluated by W. Kuhn and collaborators throughout the 1950s, who demonstrated that the osmolality of the fluids in the loop of Henle increased as they pass the renal medulla in the direction of the tip of the papilla. These authors were the first to propose the countercurrent multiplier system to explain the process by which the urine becomes more concentrated as it passes along the loop of Henle (cited in Gottschalk 1987). According to this

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model, the maximum urine concentrating ability is directly related to the length of the loops of Henle and collecting ducts that traverse the renal medulla and inversely to a nephron's diameter. Schmidt-Nielsen and O'Dell (1961) were the first to demonstrate a quantitative correlation between the relative length of the longest loops of Henle (as reflected by RMT) and the maximum urine concentration in a study that compared nine species of mammals.

Following Schmidt-Nielsen and O'Dell's (1961) seminal study, similar interspecific relationships have been shown in various groups of mammals (e.g., marsupials [Reid and McDonald 1968]; *Sylvilagus* rabbits [Heisinger and Breitenbach 1969]; cricetid rodents [Heisinger et al. 1973]; dasypodids of the Argentine desert, *Chaetophratus vellerous* and *Dasytus novemcintus* [Greegor 1975]; sciurid rodents [Blake 1977]; bats [Geluso 1978]). These studies also demonstrate a general trend for species that inhabit arid and semiarid environments to have high abilities to concentrate urine and also high RMT (see also Schmidt-Nielsen 1964; MacMillen and Lee 1967, 1969; Purohit 1974a; Borut and Shkolnik 1974).

In addition to RMT, several other morphometric indices have been proposed to estimate renal performance (Heisinger and Breitenbach 1969; Schmid 1972; Brownfield and Wunder 1976). In all cases, higher indices are typical of small desert mammals. However, RMT is the index most commonly reported in the literature (Beuchat 1996 and references cited therein). Beuchat (1996) compiled a comprehensive set of data from the literature on renal structure and function for 330 species of mammals. Among other goals, she sought to examine the influence of both body mass and habitat on kidney mass and RMT. Considering all mammals, conventional least squares linear regression analysis indicated that kidney mass scaled on body mass with a slope of 0.88, which is significantly less than unity and similar to values previously reported (Calder and Braun 1983). Interestingly, both RMT and maximum urine concentration scaled negatively with body mass (i.e., larger-bodied species had lower values), with slopes of -0.11 and -0.09 , respectively. When considering the effect of habitat (with body mass as a covariate), mammals from arid habitats tended to have greater RMT and maximum urine concentration, as compared with those from mesic and freshwater habitats. Absolute medullary thickness gave similar results when tested across habitats. Hence, allometric relationships for RMT and urine concentrating ability have been developed separately for mammals from mesic (Blake 1977; Beuchat 1996) and xeric environments (Calder and Braun 1983; Beuchat 1996).

Beuchat (1996) also reported a positive and highly statistically significant relationship between maximum urine osmolality and RMT across 78 species of mammals (her Fig. 8). However, she also reported that both traits show a negative allometric relationship with body mass (her Figs. 4, 5). Thus, the two traits might be related simply because both are correlated with body mass. This can be tested by correlating re-

siduals from the log-log regressions of each trait on body mass, but Beuchat (1996) did not report such an analysis. She did, however, report analyses for residuals of total, outer, and inner medulla thicknesses, some of which showed significant relationships with urine concentration (her Figs. 6, 7).

Aside from the issue of correcting for effects of body size, numerous studies over the past 20 years have shown that conventional statistical methods can be misleading when applied to comparative data (see, e.g., reviews in Pagel 1992; Garland et al. 1992, 1993, 1999; Garland and Ives 2000; Rohlf 2001; Rezende and Garland 2003). The existence of hierarchical phylogenetic relationships implies that data for different species cannot be considered as independent and identically distributed for purposes of statistical analyses. Ignoring phylogenetic relationships often leads to inflated Type I error rates (significant relationships are claimed too frequently) and poor estimates of parameters (e.g., slopes of allometric relationships). Moreover, studies that compare species sampled broadly with respect to phylogenetic relationships run the risk of comparing "apples and oranges" (see, e.g., Huey and Bennett 1990; Garland and Adolph 1994; Garland 2001).

The purpose of this article is to examine both RMT and kidney mass in relation to body size, habitat, and phylogeny within one clade of mammals, the Rodentia. We hypothesized that both RMT and kidney mass, corrected for body size, would correlate positively with habitat aridity and that all traits would show phylogenetic signal, that is, the tendency for related species to resemble each other (Blomberg and Garland 2002). We considered only rodents for several reasons. First, by focusing on a single lineage of mammals, we hoped to avoid comparing "apples and oranges." Second, the order Rodentia includes more than 2,000 species, which represent about half of the extant mammals (Wilson and Reeder 1993). Third, extant Rodentia vary widely in body mass, spanning approximately four orders of magnitude. Fourth, they vary widely in life-history strategies and patterns of evolutionary adaptation to different environmental settings (Eisenberg 1981). Fifth, rodents occupy a variety of habitats, and many species, representing several different evolutionary lineages (clades), inhabit arid environments. This is important because multiple evolutionary origins of a feature (e.g., occupancy of arid habitats) within a lineage increase statistical power to detect relationships with other features (e.g., RMT) in comparative studies (Garland et al. 1993; Vanhooydonck and Van Damme 1999). Sixth, more kidney data are available for Rodentia than for any other mammalian order (see Beuchat 1996). Seventh, rodents exhibit a wide range of both urine concentrating ability and RMT, which will enhance statistical power. Finally, the Rodentia has received intense scrutiny from molecular systematists in recent years, which allows construction of a reasonably well resolved estimate of phylogenetic relationships.

We greatly expand the database as compared with Beuchat (1996) by including information from several studies that were

either not included in her paper or have been published subsequently, and by presenting new data for Argentinean and Chilean rodents, emphasizing species that occur in arid habitats. In total, we consider data for kidney mass of 104 species, only 28 of which were in Beuchat's (1996) paper, and RMT of 141 species, of which 55 were in her paper. In addition, as noted above, Beuchat's (1996) demonstration of a positive relationship between maximum urine concentration and RMT is suspect because (1) she did not use regression residuals to remove negative correlations of both traits with body mass and (2) she did not account for phylogenetic relationships among species. Therefore, we reanalyze the same data that she used but correct both of these deficiencies. We also apply new methods to quantify and to test for the statistical significance of phylogenetic signal in all traits (Blomberg et al. 2003). These methods provide a guide as to whether results of conventional or phylogenetically based statistical methods should be more reliable for a given data set (see also Freckleton et al. 2002).

Material and Methods

Data Collection

We gathered new data on rodents from Argentina (E.C.-V. and C.Z.: 14 species from several arid localities) and Chile (C.Z.: 37 species from both arid and mesic localities; collecting localities are available from the authors on request). Sample size and sex ratio varied among species (see App. A), but we did not attempt to correct for this in subsequent statistical analyses because this information was often not available for the literature values (see below). For these new data, habitat assignments (see below) were based on the actual collecting localities, not general literature descriptions. Scientific names are according to Wilson and Reeder (1993) and Redford and Eisenberg (1992). Exceptions were sigmodontine species inhabiting Salta and Catamarca provinces, for which we used nomenclature recommendations of Mares et al. (1989) and Mares et al. (1997), respectively. Chilean rodents were captured from both mesic and xeric habitats. Species names follow Wilson and Reeder (1993), but *Eligmodontia*, *Phyllotis*, and *Abrothrix* are named according to Spotorno et al. (1990, 1994, 2000) and Kelt et al. (1991). We also included the new species *Laxodontomys pikumche* (Spotorno et al. 1998).

For the new data, kidneys from only adult animals and from only one locality were examined. Length and breadth of kidneys were measured with a vernier (± 0.1 mm) and sagittal half-sectioned (Cortes et al. 1990). Total width and the medullary thickness of the kidney were measured in sagittal slices from the cortex-medullary tips to the extreme of the papilla (Heisinger and Breitenbach 1969; Blake 1977) under a Wild M3 microscope. Midsagittal cuts were made to maximize the area of

visualization of the medulla. RMT was calculated following Sperber (1944):

$$\text{RMT} = \frac{10(\text{medullary thickness})}{(\text{length} \times \text{breadth} \times \text{width})^{1/3}}$$

Body mass of individuals was obtained from field records, when available; otherwise, a mean value cited in the literature was used (Redford and Eisenberg 1992; Silva and Downing 1995).

For other species (or populations), data for body mass, total mass of both kidneys, RMT, and habitat were taken from the literature. All values cited by Beuchat (1996) were checked, and the original sources are cited in Appendix A, with a few exceptions. When body mass was not reported, an estimate of average adult body mass was obtained from field guides and a variety of other sources. Most common and scientific species names follow Beuchat (1996) or Musser and Carleton (1993). All of the reported measures are for adults. Habitat was recorded as aquatic (A), mesic (M), semidesert (SD, similar to Beuchat's [1996] listing of "DM" for arid and mesic), or desert (D), which corresponds to a rank-ordering of aridity on a scale of 1–4. (For statistical analyses and figures, species listed as DM were pooled into the SD category.) We tried to obtain habitat information from the same study from which other data were collected but used other sources, including field guides, when necessary. Such categorizations can be criticized because of their crude and potentially misleading nature (Leroi et al. 1994), but this should primarily reduce statistical power to detect associations, if they exist. As will be shown, our analyses were in fact able to demonstrate associations with habitat. Clearly, an improvement for future studies would be to obtain quantitative environmental measurements from at or near the capture sites of each species (Tieleman et al. 2003; Rezende et al. 2004).

Phylogenetic Relationships

The final data set included 164 species, subspecies, or populations of rodents, ranging in body mass from 6 g to 53 kg. The overall estimate of phylogeny used in statistical analyses was derived from the literature, as detailed in Appendix B in the online edition of *Physiological and Biochemical Zoology*. Most of this phylogenetic information is from molecular-based studies, with morphological studies used when such was unavailable. When no phylogenetic information was available for a particular species, it was placed beside congeners as a polytomy, which was assumed to be "hard" (i.e., to reflect simultaneous speciation events) for purposes of phylogenetic statistical analyses (Purvis and Garland 1993; Garland and Diaz-Uriarte 1999). For each analysis, we then pruned the tree to include only those species for which data were available (maximum urine concentration, $N = 38$; kidney mass, $N = 104$; RMT, $N = 141$).

Where it was possible and warranted to keep separate studies

of a given species as separate data points (tips on the phylogeny), we did so. For example, both *Akodon albiventer* and *Abrothrix andinus* were sampled from both Argentina and Chile (new data from our study) and so were kept separate. For cases in which a given species is represented by only two “populations” in the data set, the phylogeny is obvious (i.e., a bifurcation). For three species, we had data on more than two populations for a given trait (three for *Phyllotis xanthopygus*, three for *Peromyscus leucopus*, five for *Acomys cahirinus*). In all of these cases, the samples came from geographically separate localities and so were retained as separate data points. Within each of these species, the phylogenetic relationships were treated as “hard” polytomies; thus, as in conventional statistical analyses, the separate populations were given equal weight relative to each other in phylogenetic analyses (see “Statistical Analyses”).

Because estimates of divergence times were not available for all taxa included in the data sets, we tried three different types of arbitrary branch lengths: all equal in length (constant); Grafen’s (1989), where each node is set to a depth that is one less than the number of tips that descend from it; and Pagel’s (1992), where all branches are initially set to unity and then the lengths of all branches that lead to tips (terminal taxa) are extended to make all tips contemporaneous. The adequacy of branch lengths can be checked in several ways. The most commonly used one is plotting the absolute value of the standardized phylogenetically independent contrasts versus their standard deviations and testing for a significant correlation, which would indicate that the branch lengths are inadequate (Garland et al. 1992; Garland and Díaz-Uriarte 1999). Another approach is to compare the variance of the contrasts or, in the generalized least squares mode of operation (see Blomberg et al. 2003), the mean squared error, with lower values indicating better fit of the tree to the data. Based on these two procedures, we used Pagel’s (1992) arbitrary branch lengths for all analyses, including tests for phylogenetic signal. The actual trees used, with branch lengths, are reported in standard bracket format in Appendix B and are available as PDI files (as output by the PD TREE program) from T.G.

Statistical Analyses

All traits (except habitat) were \log_{10} transformed before analyses. For completeness and to facilitate comparisons with previous studies, we used both conventional and phylogenetically based statistical analyses (reviews in Garland et al. 1999; Garland and Ives 2000; Rohlf 2001). Whether conventional or phylogenetically based results should be given greater credence depends to a large extent on whether the traits in question exhibit significant phylogenetic signal, that is, a tendency for related species to resemble each other (Blomberg and Garland 2002; see also Freckleton et al. 2002). To determine whether traits showed significant phy-

logenetic signal, we used the randomization test (1,000 permutations) implemented in the MatLab program PHYSIG.M of Blomberg et al. (2003; <http://www.biology.ucr.edu/people/faculty/Garland/PHYSIG.html>). We analyzed habitat (scored on the semiquantitative 1–4 scale), log body mass, and the logs of mass-corrected maximum urine concentration, kidney mass, and RMT. To compute mass-corrected values, we first computed the allometric scaling exponent with independent contrasts. We divided the trait (kidney mass or RMT) by body mass raised to this exponent and then took the logarithm (following Blomberg et al. 2003; an equivalent procedure is to use the “residuals” described on screen 9D of PD TREE [see below] and optionally saved in a file with extension .RSD). This procedure thus uses the phylogenetically correct estimate of the scaling relationship to adjust for correlations with body mass. We also report K (Blomberg et al. 2003), a descriptive statistic that indicates the amount of phylogenetic signal in a trait relative to the amount that would be expected for the specified phylogenetic tree (topology and branch lengths) and given a Brownian motion (random walk in continuous time) model of evolution. A K of 1 indicates that a trait has exactly the amount of signal expected, whereas values greater than 1 indicate more and values less than 1 indicate less signal than expected. The K statistic is useful for comparing the amount of signal in traits of different types (for a survey of traits from various published studies, see Blomberg et al. 2003).

Allometric equations and residuals from them were computed both the conventional way (using SPSS) and with phylogenetically independent contrasts (Felsenstein 1985; Garland et al. 1992; Garland and Adolph 1994). Assuming that the topology and branch lengths are correct, this algorithm leads to a phylogenetically independent data set consisting of $N - 1$ standardized contrasts for N original species (tips on the phylogenetic tree). Correlations and regressions with independent contrasts are computed through the origin (Felsenstein 1985; Garland et al. 1992), so degrees of freedom are the same as for conventional statistics. Independent contrasts were computed with the PD TREE module of the Phenotypic Diversity Analysis Programs (Garland et al. 1993, 1999; Garland and Ives 2000; <http://www.biology.ucr.edu/people/faculty/Garland/PDAP.html>). We computed y -intercepts and confidence intervals (CIs) following Garland and Ives (2000), as implemented in PD TREE for bivariate regressions and as implemented in REGRESSION.M (Blomberg et al. 2003) for multiple regressions.

To determine the relationship between maximum urine concentration and RMT, we used data for 38 species for which both variables were available (34 of these were in Beuchat 1996; the following values were added: *Neotoma albigula*, 2,670 mmol/kg H₂O [Brownfield and Wunder 1976]; *Gerbillurus setzeri*, 5,370, and *Gerbillurus paebe*, 4,840 [Downs and Perrin 1991; Frean et al. 1998]; *Thallamys nigricauda*, 7,630 [Frean et al. 1998]). We then computed residuals from least squares linear

regressions of each trait on body mass. We regressed (through the origin) standardized contrasts in \log_{10} urine concentration on contrasts in \log_{10} (body mass), then computed residuals, and did the same for \log_{10} RMT. We then computed the correlation (through the origin) for residual contrasts in urine concentration and RMT.

We analyzed the effect of habitat in two ways. First, we treated habitat as a categorical variable and used ANCOVA to test for the relationship between kidney mass and habitat and between RMT and habitat, with body mass as a covariate. ANCOVA was computed in the conventional way (using SPSS) and via Monte Carlo computer simulations to construct phylogenetically informed null distributions of F statistics (Garland et al. 1993; PDAP modules PDSINGLE, PDSIMUL, PDANOVA). Using biologically realistic ranges for body mass, RMT, and kidney mass (on the \log_{10} scales), we performed 1,000 simulations under a gradual Brownian motion model (Felsenstein 1985) but with limits to character evolution. Following most previous uses of PDSIMUL (e.g., Garland et al. 1993; Cruz-Neto et al. 2001; Hutcheon et al. 2002), we used the Replace option to implement limits. For body mass, we used a lower limit of 2 g and an upper limit of 100 kg. The former is slightly lower than the smallest body mass in our data, Arizona pocket mouse (*Perognathus amplus*), and the latter is roughly twice the mass of the largest extant rodent, the Capybara (*Hydrochaeris hydrochaeris*). We used a lower limit of 1 and an upper limit of 20 for RMT. The former is the smallest possible value of RMT, and the latter is slightly higher than the greatest RMT value in our data set, 17.6 for the desert pocket mouse (*Chaetodipus penicillatus*). The lower limit used for kidney mass was 0.05 g, and the upper limit was 100 g. The lower limit is below the smallest kidney mass in our data set (0.08 g for *Mus musculus*), and the upper limit is above the largest kidney mass (69.75 g for *Hydrochaeris hydrochaeris*). Starting values for all simulations were the defaults in PDSIMUL, which are the conventional means of the characters. The correlation between the two traits (\log [body mass] and either \log RMT or \log [kidney mass]) was set to 0 so that we could test the null hypothesis of no effect of body mass.

Second, we treated habitat as a semiquantitative variable. Here, we performed multiple regressions (either conventional or with phylogenetically independent contrasts) of kidney mass or RMT on both body mass and habitat, and we report the one-tailed P value for habitat because we had the directional hypothesis that animals from more arid habitats would have larger kidneys and/or higher RMT after controlling for effects of body mass. This latter approach should increase statistical power to detect an effect of habitat (see, e.g., discussion in Garland et al. 1993) unless the 1–4 scale is very imprecise and/or far from being a linear approximation of variation in selective regime that is caused by variation in habitat aridity.

Results

Phylogenetic Signal

All traits showed statistically significant phylogenetic signal. For the data set that included body mass, RMT, and maximum urine concentration ($N = 38$, all \log_{10} transformed), P values were <0.001 , 0.002 , and 0.008 , respectively, with corresponding K values of 0.658 , 0.452 , and 0.362 . Signal was also present for \log_{10} of mass-corrected RMT ($P = 0.005$, $K = 0.358$) and maximum urine concentration ($P = 0.024$, $K = 0.331$).

For the 141 species in the RMT data set, body mass, RMT, and mass-corrected RMT (all \log_{10}) all showed signal at $P < 0.001$, and K values were 0.678 , 0.302 , and 0.192 , respectively. For the 104 species in the kidney mass data set, again all traits showed signal at $P < 0.001$, with K values of 0.551 for \log_{10} (body mass), 0.485 for \log_{10} (kidney mass), and 0.200 for \log_{10} (mass-corrected kidney mass).

Finally, habitat (scored on the semiquantitative 1–4 scale) showed highly significant signal ($P < 0.001$) for both the 141- and 104-species data sets, with K values of 0.344 and 0.279 , respectively. The presence of significant phylogenetic signal in all traits offers justification for the use of phylogenetically based statistical methods and suggests that results from those methods should be more reliable than those from conventional analyses.

Body Mass in Relation to Habitat

Considering the 141 species for which RMT data were available, conventional ANOVA indicated a highly significant difference among habitats ($F_{3,140} = 16.91$, $P < 0.00005$), and this was also significant in comparison with the F values from phylogenetically simulated data (critical $F = 6.46$, $P < 0.001$). As can be seen in Figure 1, mean \log_{10} (body masses) (kg) of the aquatic species (0.625 [95% confidence interval = -0.507 to 1.757], $N = 5$) averaged much greater than those of the mesic (-0.907 [-1.133 to -0.682], $N = 40$), semidesert (-1.310 [-1.591 to -1.030], $N = 22$), or desert species (-1.262 [-1.387 to -1.138], $N = 74$), and Scheffé's multiple range comparison indicated that aquatic species are significantly larger than all other groups.

Treating habitat as a quantitative variable, a conventional regression indicated a highly significant prediction of \log (body mass) ($r^2 = 0.147$, $F_{1,139} = 24.0$, two-tailed $P < 0.00005$), and independent contrasts analysis also indicated a significant relationship ($r^2 = 0.047$, $F_{1,139} = 6.83$, two-tailed $P = 0.0100$). Considering only the 90 Muridae (the only family that included all four habitat types), the conventional regression still indicated a significant prediction of \log body mass ($r^2 = 0.045$, $F_{1,88} = 4.17$, two-tailed $P = 0.0442$), but the independent contrasts analysis did not ($r^2 = 0.029$, $F_{1,88} = 2.61$, two-tailed $P = 0.1099$).

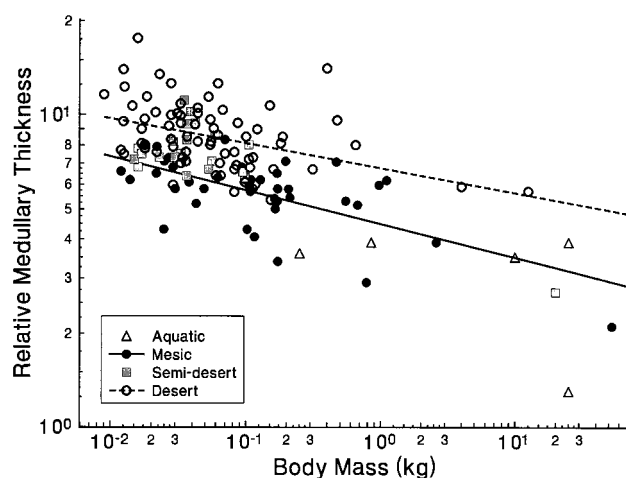


Figure 1. Relative medullary thickness (RMT) in relation to habitat for 141 species or populations of rodents. Conventional least squares linear regressions (fitted to the log-transformed data) are shown as a heuristic for both mesic and desert species. Both conventional and phylogenetically informed ANCOVA indicate significant differences among habitat types in mass-corrected RMT (see text).

Correlation between RMT and Maximum Urine Concentration

For the subset of 38 species with data for both traits, both urine concentration (UC) and RMT showed a statistically significant negative scaling with body mass. The conventional allometric equations were (95% confidence intervals are given in parentheses; body mass [M_b] units are kilograms):

$$\log_{10} \text{UC (mmol/kg H}_2\text{O)} = 3.29(3.16, 3.42) \\ - 0.301(-0.404, -0.197) \times \log_{10} M_b$$

($r^2 = 0.493$, $F_{1,36} = 35.0$, $P < 0.00005$) and

$$\log_{10} \text{RMT} = 0.663(0.572, 0.754) \\ - 0.185(-0.258, -0.111) \times \log_{10} M_b$$

($r^2 = 0.420$, $F_{1,36} = 26.0$, $P < 0.00005$). Residuals from these regression equations showed a highly significant positive relationship ($r = 0.649$, two-tailed $P < 0.00005$).

The independent contrasts allometric equations were similar in terms of slopes and intercepts, although, as expected (see Garland and Ives 2000), the 95% confidence intervals were wider:

$$\log_{10} \text{UC (mmol/kg H}_2\text{O)} = 3.24(2.82, 3.66) \\ - 0.368(-0.514, -0.222) \times \log_{10} M_b$$

($r^2 = 0.421$, $F_{1,36} = 26.2$, $P < 0.00005$) and

$$\log_{10} \text{RMT} = 0.669(0.388, 0.950) \\ - 0.180(-0.277, -0.082) \times \log_{10} M_b$$

($r^2 = 0.280$, $F_{1,36} = 14.0$, $P = 0.0006$). Again, residuals from these independent contrasts allometric equations showed a highly significant positive relationship ($r = 0.685$, $P < 0.00005$). Thus, both conventional and phylogenetic analyses validate the use of RMT as an indicator of interspecific variation in maximum urine concentrating ability in rodents.

RMT in Relation to Body Mass and Habitat

Considering all 141 data points, the conventional allometric equation was (95% confidence intervals are given in parentheses):

$$\log_{10} \text{RMT} = 0.612(0.658, 0.730) \\ - 0.145(-0.172, -0.118) \times \log_{10} M_b$$

($r^2 = 0.440$, $F_{1,139} = 109.4$, $P < 0.00005$). The independent contrasts allometric equation was similar in terms of slope and intercept:

$$\log_{10} \text{RMT} = 0.699(0.473, 0.925) \\ - 0.132(-0.180, -0.083) \times \log_{10} M_b$$

($r^2 = 0.172$, $F_{1,139} = 28.8$, $P < 0.00005$).

In the conventional ANCOVA of habitat groups, $\log M_b$ had a highly significant ($P < 0.001$) negative effect on $\log \text{RMT}$ (pooled within-groups slope = -0.102), which also differed significantly among habitats ($P < 0.001$), with desert rodents having the highest RMT values, followed by semidesert rodents, then mesic rodents, and finally aquatic animals (Fig. 1; Table 1). Critical F values obtained from the Monte Carlo simulations were, as expected, higher than conventional values, but the effects of both body mass and habitat were still significant (Table 1). The interaction between \log body mass and habitat was not significant ($P = 0.63$).

The multiple regression approach, which treated habitat as a quantitative variable, also indicates that it was a significant predictor of RMT. The conventional multiple regression yielded

Table 1: ANCOVA comparing log RMT for 141 rodents from different habitats (arid, semiarid, mesic, aquatic) with log body mass (kg) as the covariate

Source of Variation	Sum of Squares	df	Mean Square	F	Conventional		Phylogenetic	
					Critical Value	P	Critical Value	P
Habitat	.616	3	.205	21.70	2.67	<.001	5.75	<.001
log(body mass)	.544	1	.544	57.44	3.91	<.001	40.65	.025
Explained	2.114	4	.529	55.82	2.44	<.001	12.74	<.001
Error	1.288	136	.009					
Total	3.402	140	.024					

the following predictive equation (habitat is scored on a 1–4 scale):

$$\begin{aligned} \log_{10} \text{RMT} &= 0.500(0.445, 0.556) \\ &- 0.107(-0.131, -0.082) \times \log_{10} M_b \\ &\quad (\text{one-tailed } P < 0.00005) \\ &+ 0.074(0.056, 0.092) \times \text{habitat} \\ &\quad (\text{one-tailed } P < 0.00005) \end{aligned}$$

(multiple $r^2 = 0.619$, $F_{2,138} = 112.3$, $P < 0.00005$). The independent contrasts equation was

$$\begin{aligned} \log_{10} \text{RMT} &= 0.478 (\text{SE} = 0.1051434) \\ &- 0.100(-0.143, -0.057) \times \log_{10} M_b \\ &\quad (\text{one-tailed } P < 0.00005) \\ &+ 0.082(0.058, 0.106) \times \text{habitat} \\ &\quad (\text{one-tailed } P < 0.00005) \end{aligned}$$

(multiple $r^2 = 0.374$, $F_{2,138} = 41.3$, $P < 0.00005$). Repeating these analyses for the 90 Muridae only yielded a conventional multiple regression of

$$\begin{aligned} \log_{10} \text{RMT} &= 0.470(0.384, 0.557) \\ &- 0.103(-0.153, -0.054) \times \log_{10} M_b \\ &\quad (\text{one-tailed } P = 0.0001) \\ &+ 0.080(0.059, 0.100) \times \text{habitat} \\ &\quad (\text{one-tailed } P < 0.00005) \end{aligned}$$

(multiple $r^2 = 0.523$, $F_{2,87} = 47.8$, $P < 0.00005$). The independent contrasts equation was

$$\begin{aligned} \log_{10} \text{RMT} &= 0.516 (\text{SE} = 0.085537) \\ &- 0.056(-0.121, 0.001) \times \log_{10} M_b \\ &\quad (\text{one-tailed } P = 0.0475) \\ &+ 0.088(0.065, 0.111) \times \text{habitat} \\ &\quad (\text{one-tailed } P < 0.00005) \end{aligned}$$

(multiple $r^2 = 0.431$, $F_{2,87} = 32.9$, $P < 0.00005$).

Kidney Mass in Relation to Body Mass and Habitat

Considering all 104 data points, the conventional allometric equation was (95% confidence intervals are given in parentheses):

$$\begin{aligned} \log_{10}(\text{kidney mass}) \text{ (g)} &= 0.723(0.646, 0.799) \\ &+ 0.809(0.751, 0.868) \times \log_{10} M_b \end{aligned}$$

($r^2 = 0.881$, $F_{1,102} = 756.0$, $P < 0.00005$). Parameters of the independent contrasts allometric equation were similar:

$$\begin{aligned} \log_{10}(\text{kidney mass}) \text{ (g)} &= 0.680(0.310, 1.049) \\ &+ 0.773(0.676, 0.871) \times \log_{10} M_b \end{aligned}$$

($r^2 = 0.710$, $F_{1,102} = 249.4$, $P < 0.00005$).

In the conventional ANCOVA of habitat groups (Table 2), $\log M_b$ had a highly significant ($P < 0.001$) positive effect on $\log(\text{kidney mass})$ (pooled within-groups slope = -0.803), but habitat had no effect ($P = 0.318$; see Fig. 2). Comparison of F ratios with those obtained from the Monte Carlo simulations leads to similar conclusions (Table 2; the interaction between $\log M_b$ and habitat was not significant, $P = 0.89$).

Table 2: ANCOVA comparing log(kidney mass) (g) for 103 rodents from different habitats (arid, semiarid, mesic) with log(body mass) (kg) as the covariate

Source of Variation	Sum of Squares	df	Mean Square	F	Conventional		Phylogenetic	
					Critical Value	P	Critical Value	P
Habitat	.075	2	.038	1.16	3.09	.318	8.27	.664
log(body mass)	21.198	1	21.198	652.76	3.94	<.001	32.03	<.001
Explained	23.573	3	7.858	241.96	2.70	<.001	14.89	<.001
Error	3.215	99	.032					
Total	26.788	102	.263					

Note. Only one species was categorized as aquatic, so this category and data point were excluded.

In the conventional multiple regression, habitat was not a significant predictor of kidney mass:

$$\log_{10}(\text{kidney mass}) \text{ (g)} = 0.698(0.572, 0.824) + 0.814(0.752, 0.876) \times \log_{10} M_b$$

(one-tailed $P < 0.00005$)

$$+ 0.010(-0.030, 0.050) \times \text{habitat}$$

(one-tailed $P = 0.3107$)

(multiple $r^2 = 0.881$, $F_{2,101} = 375.3$, $P < 0.00005$).

However, with phylogenetically independent contrasts, both body mass and habitat were significant:

$$\log_{10}(\text{kidney mass}) \text{ (g)} = 0.575 \text{ (SE} = 0.192949) + 0.793(0.695, 0.892) \times \log_{10} M_b$$

(one-tailed $P < 0.00005$)

$$+ 0.042(-0.003, 0.088) \times \text{habitat}$$

(one-tailed $P = 0.0346$)

(multiple $r^2 = 0.719$, $F_{2,101} = 129.3$, $P < 0.00005$). When the analysis was repeated for the 71 Muridae only, the conventional multiple regression was

$$\log_{10}(\text{kidney mass}) \text{ (g)} = 0.848(0.604, 1.092) + 0.870(0.734, 1.007) \times \log_{10} M_b$$

(one-tailed $P < 0.00005$)

$$- 0.013(-0.066, 0.040) \times \text{habitat}$$

(two-tailed $P = 0.6152$, sign in wrong direction)

(multiple $r^2 = 0.708$, $F_{2,68} = 82.5$, $P < 0.00005$). The independent contrasts equation was

$$\log_{10}(\text{kidney mass}) \text{ (g)} = 0.708 \text{ (SE} = 0.21132) + 0.890(0.733, 1.048) \times \log_{10} M_b$$

(one-tailed $P < 0.00005$)

$$+ 0.033(-0.027, 0.093) \times \text{habitat}$$

(one-tailed $P = 0.1378$)

(multiple $r^2 = 0.654$, $F_{2,68} = 64.2$, $P < 0.00005$).

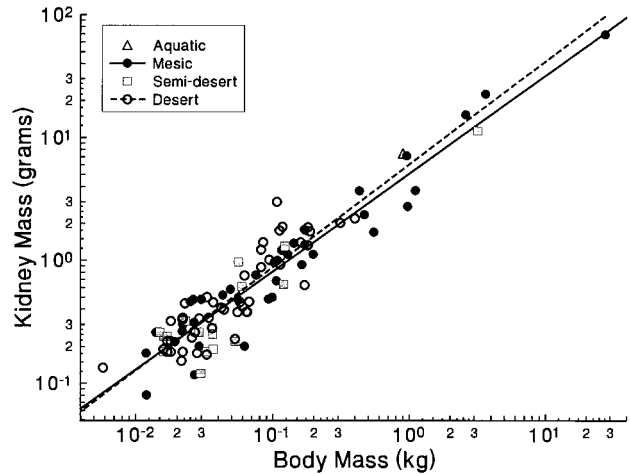


Figure 2. Kidney mass in relation to habitat for 104 species or populations of rodents. Conventional least squares linear regressions (fitted to the log-transformed data) are shown as a heuristic for both mesic and desert species. Both conventional and phylogenetically informed ANCOVA indicate no significant differences among habitat types in mass-corrected kidney mass (see text). However, a phylogenetic analysis that treated habitat as a quantitative variable (scored on a 1–4 scale to indicate increasing aridity) revealed a significant positive relationship between mass-corrected kidney mass and habitat (see text).

Discussion

This is the first study to employ an explicitly phylogenetic analysis to examine the effects of body size and of habitat on the most commonly used morphological indicator of mammalian kidney performance, relative medullary thickness (RMT), as well as kidney mass. For reasons presented in the "Introduction," analyses were restricted to a single mammalian lineage, the Rodentia. Phylogenetic signal was statistically significant for all traits examined (including habitat), which suggests that results of phylogenetically informed statistical analyses should be more reliable than conventional analyses (Freckleton et al. 2002; Blomberg et al. 2003). Nevertheless, results of conventional and phylogenetic analyses were largely congruent, with one interesting exception (see below).

Although body mass showed significant phylogenetic signal, as has been reported previously for rodents and other animal groups (Freckleton et al. 2002; Blomberg et al. 2003; Rezende et al. 2004), rodents categorized as aquatic (nutria, beaver, muskrat, water rat) were significantly larger in body mass as compared with species from all other habitat types. Thus, both ecology and phylogeny affect body size. As body size affects virtually all aspects of an organism, various possible adaptive explanations for the large size of aquatic rodents can be hypothesized. For example, natural selection in aquatic habitats might favor large size because it lowers surface/volume ratios, which in turn confer lower mass-specific rates of heat loss. Another possibility is that aquatic habitats tend to offer higher primary productivity, which leads to selection that favors large body size. Alternatively, an aquatic lifestyle may reduce constraints on body size that relate to locomotor biomechanics.

After correcting for correlations with body mass, RMT was a highly significant predictor of maximum urine concentrating ability, thus bolstering its use as a comparative indicator of mammalian kidney performance. Consistent with previous studies, we found that RMT scaled negatively with body mass and that kidney mass scaled with an exponent significantly less than unity, such that larger-bodied species of rodents generally have lower values for RMT and relatively smaller kidneys. We also found that RMT (corrected for body size) varies significantly among habitats, with species from arid habitats having higher RMT, as would be expected since the pioneering work of Sperber (1944; see "Introduction").

Kidney mass (corrected for body size) tended to increase with increasing habitat aridity (scored on a semiquantitative scale), which was statistically significant in the phylogenetic analysis (one-tailed $P = 0.0346$) but not in the conventional analysis (one-tailed $P = 0.3107$). Because the P value for habitat differed greatly between conventional and phylogenetic analyses, we checked for outliers and/or influential points in both (e.g., magnitude of standardized residual, leverage). The conventional regression did not yield any noteworthy residuals (e.g., all standardized residuals were <3.0 in magnitude). The

largest standardized residual was 2.92 for *Parotomys littledalei*. The leverage for this data point was not unusually large (0.0141), and the outlier test described in Cook and Weisberg (1999) produced a P value of 0.28. Similarly, the independent contrasts regression yielded no standardized residuals >3.0 . The largest value was -2.96 for node 63 on the tree, which contrasts *Peromyscus eremicus* (habitat D) with the ancestor of two populations of *Peromyscus leucopus* (one from South Dakota with SD habitat, the other from Pennsylvania with M habitat). The leverage for this point was only 0.0289, and the outlier test indicated $P = 0.24$. Still, the large difference in relative kidney mass between *P. eremicus* (0.98% of body mass) and the average of the two *P. leucopus* populations (0.42%) is noteworthy and suggests that this genus, which shows substantial physiological variability (MacMillen and Garland 1989), would be a good candidate for common-garden comparative studies (Garland and Adolph 1991, 1994; Oswald 1998). This example also demonstrates that phylogenetic analyses can sometimes uncover relationships (yield greater statistical power) that are not evident when phylogeny is ignored.

Comparative phylogenetic analyses seek to elucidate how traits have evolved and are predicated on the assumption that observed differences among species are at least mostly genetically based. However, all of the traits considered here (body mass, kidney mass, RMT) can be affected by environmental conditions experienced from birth (or even before) into adulthood (Bankir et al. 1988; Garland and Adolph 1991; Oswald 1998; Al-kahtani 2003; references therein), and with very few exceptions the data analyzed herein were from wild-caught individuals that must have experienced sometimes immensely different environmental conditions during their ontogeny. Thus, it is possible that a substantial fraction of the interspecific variation that we analyzed could represent direct environmental effects rather than genetically based differences. For example, Al-kahtani (2003) found that chronic water restriction applied to outbred house mice (Hsd:ICR strain) from weaning at 21 d of age reduced body mass by 50% at 146 d of age (mass of free-water group was 36.14 ± 0.589 [mean \pm SE] g, $N = 15$; mass of water-restricted group was 18.21 ± 0.524 g, $N = 19$). For these same individuals, body mass-adjusted kidney mass (from ANCOVA) was significantly increased by 12% (free-water group averaged 0.459 ± 0.008 g vs. 0.512 ± 0.009 g), but RMT was not significantly affected (raw means averaged 0.729 ± 0.028 [$N = 14$] for the free-water group and 0.726 ± 0.025 [$N = 15$] for the water-restricted group). Although many factors other than water availability can also affect these traits, and it is possible that laboratory house mice are less plastic than some wild rodents, Al-kahtani's (2003) results suggest that the differences among species analyzed in our study cannot represent entirely environmental differences experienced during ontogeny. For example, body mass varies by 9,000-fold in our sample, relative kidney mass varies by more than twofold even among species of *Peromyscus* (see previous paragraph), and

inspection of Appendix A indicates many cases of substantial variation in RMT among species within a single genus (e.g., 26% among species of *Octodon*, 33% among species of *Eutamias*, 68% among species of *Meriones*, 30% between two species of *Cynomys*, 82% between two species of *Chaetodipus*).

Many previous studies have shown that RMT varies inversely with body size in different groups of mammals (Blake 1977; Greigor 1975; Geluso 1978; Beuchat 1991, 1993, 1996). For example, Beuchat (1991) examined the relationship between body mass and RMT for a taxonomically diverse sample of 165 species of mammals. She reported that both RMT and urine concentrating ability scaled negatively with body mass as: $RMT = 5.408M^{-0.108}$ and $U_{osm} = 2,564M^{-0.097}$, respectively. Thus, small mammals have higher RMT and produce more concentrated urine as compared with larger-bodied mammals. Our analyses demonstrate that these conclusions hold for rodents and also that RMT is a significant predictor of maximum urine concentration after controlling statistically for effects of both body size and phylogeny.

It seems a paradox that larger mammals with their absolutely longer loops of Henle are unable to concentrate their urine to the same degree as most small mammals. This pattern is inconsistent with the assumption of a countercurrent multiplier system, which states that the longer loops facilitate the production of more concentrated urine. Greenwald and colleagues (Greenwald and Stetson 1988; Greenwald 1989; Abrahams et al. 1991) have recognized this discrepancy and offered a “metabolic hypothesis” to explain it. This hypothesis is based on the premise that small mammals have a higher mass-specific metabolic rate than larger ones, and hence their absolutely shorter loops are metabolically more active (per unit tissue) than those of large mammals. They have tested this hypothesis with data from several mammals ranging in body mass from a horse of 400 kg to a bat of 11 g. They concluded that loops of Henle, especially the medullary thick ascending limb (MTAL) of small mammals have (1) more infoldings in the basolateral membrane per unit volume, (2) more mitochondria per unit volume, and (3) more inner mitochondrial membrane per unit volume. These structural modifications are crucial elements in the metabolic rate of ATP production, which drive the active transport of NaCl in this segment and hence maintain a higher cortical-medullary urine concentration gradient.

In our analysis of rodents, body mass–corrected RMT was positively associated with habitat aridity. This result is consistent with the finding of convergent trends in kidney structure, including a thick medulla relative to kidney size, that have been observed in a variety of lineages of small mammals from diverse desert habitats (e.g., rabbits [Heisinger and Breitenbach 1969], hedgehogs [Yaakobi and Shkolnik 1974], insectivorous bats [Geluso 1978], and heteromyid rodents [MacMillen and Hinds 1983]). However, as noted in the previous paragraph, maximum loop length alone does not determine variation in urine concentrating ability. Other important structural characteristics include the arrangement of vascular bundles (the vasa recta) within the medulla, nephron heterogeneity, the presence of extensions of the renal pelvis into the medulla (specialized pelvic fornices), and the degree of confluence of the collecting ducts in the inner medulla (Dantzler and Braun 1980; Bankir and de Rouffignac 1985; Braun 1998). When data for these traits become available for a greater range of species, it will be of interest to perform analyses similar to what we have done for RMT and kidney mass. It will also be of interest to examine other morphometric indices (e.g., see Brownfield and Wunder 1976) and molecular indicators of kidney function, such as vasopressin binding (Oswald 1998; Al-kahtani 2003).

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Appendix A

Table A1: Body mass, total mass of both kidneys ($N = 104$), relative medullary thickness ($N = 141$), and habitat of rodents

Tip Code	Tip No.	Genus and Species (Common Name)	Family	RMT	Kidney Mass (g)	Body Mass (kg)	Habitat	Reference
HC	1	<i>Hystrix cristata</i> (African porcupine)	Hystriidae	2.7		20	DM	Sperber 1944
DP	2	<i>Dolichotis panagonica</i> (Patagonian cavy)	Caviidae	5.7		12.5	D	Sperber 1944
Mn	3	<i>Microcavia niata</i> (Chile; 4, 3)	Caviidae	8.1	1.86	.181	D	This study (C.Z.)
CZ	4	<i>Cavia porcellus</i> (guinea pig)	Caviidae		3.698	.43	M	Spector 1956
Ap	5	<i>Agouti</i> (= <i>Cuniculus paca virgatus</i>) (spotted agouti)	Dasyproctidae		22.7	3.627	M	Quiring 1950
DI	6	<i>Dasyprocta leporina</i> (= <i>aguti</i>) (agouti)	Dasyproctidae	3.9	15.39	2.6	M	Sperber 1944
Hh	7	<i>Hydrochaeris hydrochaeris</i> (= <i>isthimius</i>) (capybara)	Hydrochoeridae	2.1		53	M	Sperber 1944
HH	8	<i>H. hydrochaeris</i> (= <i>isthimius</i>) (capybara)	Hydrochoeridae		69.75	27.67	M	Sperber 1944; Quiring 1950
aC	9	<i>Abrocoma cinerea</i> (Chile; 0, 4)	Abrocomidae	6.9	1.40	.085	D	This study (C.Z.)
aB	10	<i>Abrocoma bennetti</i> (Chile; 3, 2)	Abrocomidae	7.1	1.12	.197	M	This study (C.Z.)
cL	11	<i>Chinchilla lanigera</i> (Chile; 4, 2)	Chinchillidae	6.7	2.02	.312	D	This study (C.Z.)
Cl	12	<i>Chinchilla laniger</i>	Chinchillidae	9.6		.475	D	Sperber 1944; Weisser et al. 1970
My	13	<i>Myocastor coypus</i> (nutria)	Capromyidae	3.5		10	A	Pfeiffer 1970; Sperber 1944
Ce	14	<i>Ctenomys eremophilus</i>	Octodontoidae	8.98		.12118	D	Diaz and Ojeda 1999
O2	15	<i>Octodontomys gliroides</i> [sic]	Octodontoidae	5.35		.153	D	Diaz and Ojeda 1999
Og	16	<i>Octodontomys glyroides</i> (Chile; 2, 3)	Octodontoidae	8.5	1.72	.187	D	This study (C.Z.)
Tb	17	<i>Tympanoctomys barrerae</i>	Octodontoidae	9.41		.08679	D	Diaz and Ojeda 1999
Om	18	<i>Octomys mimax</i>	Octodontoidae	6.09		.09832	D	Diaz and Ojeda 1999
oD	19	<i>Octodon degus</i> (Chile; 7, 2)	Octodontoidae	6.7	1.40	.16	D	This study (C.Z.)
oB	20	<i>Octodon bridgesi</i> (Chile; 5, 3)	Octodontoidae	5.4	.92	.163	M	This study (C.Z.)
oL	21	<i>Octodon lunatus</i> (Chile; 2, 0)	Octodontoidae	5.3	1.34	.171	M	This study (C.Z.)
Af	22	<i>Aconaemys fuscus</i> (Chile; 2, 1)	Octodontoidae	6.2	1.12	.128	M	This study (C.Z.)
As	23	<i>Aconaemys sagei</i> (Chile; 3, 0)	Octodontoidae	5.9	.68	.106	M	This study (C.Z.)
S2	24	<i>Spalacopus cyanus</i> (Chile; 5, 4)	Octodontoidae	6.2	.50	.099	M	This study (C.Z.)
Ci	25	<i>Castor fiber</i> (European beaver)	Castoridae	3.9		25	A	Sperber 1944
Cc	26	<i>Castor canadensis</i> (American beaver)	Castoridae	1.3		25	A	Schmidt-Nielsen and O'Dell 1961; Munkacs and Palkovits 1977
Cv	27	<i>Ctenodactylus vali</i> (gundi)	Ctenodactylidae		1.328	.18	D	de Rouffignac et al. 1981
AR	28	<i>Aplodontia rufa</i> (mountain beaver)	Aplodontoidae	2.9		.785	M	Dolph et al. 1962; Nungesser and Pfeiffer 1965; Sperber 1944
Ts	29	<i>Tamias striatus</i> (eastern chipmunk)	Sciuridae	6.2		.109	M	Blake 1977
TS	30	<i>T. striatus</i> (eastern chipmunk) second study	Sciuridae		.756	.075	M	Quiring 1950

Table A1 (Continued)

Tip Code	Tip No.	Genus and Species (Common Name)	Family	RMT	Kidney Mass (g)	Body Mass (kg)	Habitat	Reference
Es	31	<i>Eutamias speciosus</i> (lodgepole chipmunk)	Sciuridae	8.32		.07	M	Heller and Poulson 1972
Em	32	<i>Eutamias minimus</i> (least chipmunk)	Sciuridae	11.1		.035	DM	Heller and Poulson 1972
EA	33	<i>Eutamias amoenus</i> (yellow-pine chipmunk)	Sciuridae	9.39		.042	DM	Heller and Poulson 1972
Ea	34	<i>Eutamias alpinus</i> (alpine chipmunk)	Sciuridae	10.2		.039	DM	Heller and Poulson 1972
Xi	35	<i>Xerus inauris</i> (African ground squirrel)	Sciuridae	14.1	2.2	.4	D	Marsh et al. 1978
Fp	36	<i>Funambulus pennanti</i> (five-striped squirrel)	Sciuridae	6.8	.484	.0929	M	Purohit et al. 1973; Purohit and Ghosh 1965
Sp	37	<i>Spermophilus parryii parryii</i> (arctic ground squirrel)	Sciuridae		7.12	.958	M	Quiring 1950
Sb	38	<i>Spermophilus beecheyi</i> (California ground squirrel)	Sciuridae	7.07	2.36	.468	M	Baudinette 1974
Sl	39	<i>Spermophilus lateralis</i> (golden-mantled ground squirrel)	Sciuridae	5.44		.212	M	Blake 1977; Munkacsi and Palkovits 1977
Cu	40	<i>Cynomys ludovicianus</i> (black-tailed prairie dog)	Sciuridae	5.96	2.76	.972	M	Harlow and Braun 1995
cl	41	<i>Cynomys leucurus</i> (white-tailed prairie dog)	Sciuridae	6.17	3.72	1.11	M	Harlow and Braun 1995
TH	42	<i>Tamiasciurus hudsonicus</i> (red squirrel)	Sciuridae	5.79		.207	M	Bakko 1975
Th	43	<i>T. hudsonicus</i> (red squirrel)	Sciuridae		1.3815	.1425	M	Layne 1954
Sv	44	<i>Sciurus vulgaris</i> (tree squirrel)	Sciuridae	5.3	1.7	.55	M	Sperber 1944
Sc	45	<i>Sciurus carolinensis</i> (gray squirrel)	Sciuridae	5.15		.673	M	Bakko 1975
Pc	46	<i>Pedetes capensis</i> (= <i>caffer</i>) (spinghare)	Pedetidae	5.9		4	D	Sperber 1944
P2	47	<i>P. capensis</i> (springhare)	Pedetidae		11.37	3.18	SD	Butynski 1979
DG	48	<i>Geomys pinetis</i>	Geomyidae	5		.165	M	Sperber 1944; Hickman and Brown 1973
Mp	49	<i>Microdipodops pallidus</i> (pale kangaroo mouse)	Heteromyidae	9.52		.0125	D	Lawler and Geluso 1986
Dd	50	<i>Dipodomys deserti</i> (desert kangaroo rat)	Heteromyidae	7.21		.105	D	Lawler and Geluso 1986
Dm	51	<i>Dipodomys merriami</i> (Merriam's kangaroo rat)	Heteromyidae	8.5	.4534	.0367	D	Sperber 1944; Carpenter 1966; Munkacsi and Palkovits 1977; Altschuler et al. 1979
Do	52	<i>Dipodomys ordii</i> (Ord's kangaroo rat)	Heteromyidae	8.2		.044	D	Lawler and Geluso 1986
Ds	53	<i>Dipodomys spectabilis</i> (banner-tailed kangaroo rat)	Heteromyidae	8.5		.1	D	Schmidt-Nielsen et al. 1948; Munkacsi and Palkovits 1977
DA	54	<i>Dipodomys agilis</i>	Heteromyidae	8		.0543	D	Sperber 1944; Price and Longland 1989
DM	55	<i>Dipodomys microps</i>	Heteromyidae	8.61	.75	.062	D	Breyen et al. 1973; Ojeda et al. 1999

Table A1 (Continued)

Tip Code	Tip No.	Genus and Species (Common Name)	Family	RMT	Kidney Mass (g)	Body Mass (kg)	Habitat	Reference
Cf	56	<i>Chaetodipus formosus</i> (long-tailed pocket mouse)	Heteromyidae	9.68		.0179	D	Lawler and Geluso 1986
CP	57	<i>Chaetodipus penicillatus</i> (desert pocket mouse)	Heteromyidae	17.6	.1886	.0159	D	Munkacsi and Palkovits 1977; Altschuler et al. 1979; Beuchat 1996
Cb	58	<i>Chaetodipus baileyi</i> (Bailey's pocket mouse)	Heteromyidae		.2352	.0257	D	Altschuler et al. 1979
Pa	59	<i>Perognathus amplus</i> (Arizona pocket mouse)	Heteromyidae		.1338	.0058	D	Altschuler et al. 1979
Pl	60	<i>Perognathus longimembris</i> (little pocket mouse)	Heteromyidae	11.6		.009	D	Beuchat 1996
Zh	61	<i>Zapus hudsonius</i> (meadow jumping mouse)	Zapodidae		.218	.0193	M	Quiring 1950
Jj	62	<i>Jaculus jaculus</i> (jerboa)	Dipodidae	9.3	.41	.042	D	Sperber 1944; Schmidt-Nielsen et al. 1948; Munkacsi and Palkovits 1965
Ne	63	<i>Neotomys ebriosus</i> (Chile; 2, 0)	Muridae	6.3	.20	.062	M	This study (C.Z.)
Ol	64	<i>Oligoryzomys longicaudatus</i> (Chile; 5, 4)	Muridae	6.8	.20	.029	M	This study (C.Z.)
Ax	65	<i>Abrothrix xanthorhinus</i> (Chile; 4, 1)	Muridae	7.9	.30	.022	M	This study (C.Z.)
Al	66	<i>Abrothrix longipilis</i> (Chile; 6, 3)	Muridae	7.1	.46	.025	M	This study (C.Z.)
Ao	67	<i>Abrothrix olivaceus</i> (Chile; 3, 4)	Muridae	8.0	.22	.018	D	This study (C.Z.)
Aa	68	<i>Abrothrix andinus</i> (Chile; 4, 2)	Muridae	7.6	.18	.022	D	This study (C.Z.)
aa	69	<i>A. andinus</i> (Argentina; 6, 3)	Muridae	8.1	.22	.017	D	This study (E.C.-V. and C.Z.)
CM	70	<i>Chelemys macronyx</i> (Chile; 4, 1)	Muridae	5.8	.58	.049	M	This study (C.Z.)
Bl	71	<i>Bolomys lactens</i> (Argentina; 2, 0)	Muridae	7.1	.28	.036	SD	This study (E.C.-V. and C.Z.)
Av	72	<i>Akodon varius</i> (Argentina; 5, 2)	Muridae	7.5	.18	.033	SD	This study (E.C.-V. and C.Z.)
AB	73	<i>Akodon berlepschii</i> (Chile; 3, 1)	Muridae	7.8	.32	.018	D	This study (C.Z.)
AA	74	<i>Akodon albiventer</i> (Chile; 1, 4)	Muridae	7.6	.34	.022	D	This study (C.Z.)
aA	75	<i>A. albiventer</i> (Argentina; 4, 4)	Muridae	7.3	.32	.023	SD	This study (E.C.-V. and C.Z.)
am	76	<i>Akodon molinae</i>	Muridae	10.09		.03111	D	Diaz and Ojeda 1999
cc	77	<i>Calomys callosus</i> (Argentina; 8, 1)	Muridae	6.8	.24	.016	SD	This study (E.C.-V. and C.Z.)
CL	78	<i>Calomys lepidus</i> (Chile; 4, 0)	Muridae	6.2	.26	.014	M	This study (C.Z.)
Cm	79	<i>Calomys musculinus</i>	Muridae	12.29		.01273	D	Diaz and Ojeda 1999
cm	80	<i>C. musculinus</i> (Argentina; 5, 3)	Muridae	7.2	.26	.015	SD	This study (E.C.-V. and C.Z.)
Ep	81	<i>Eligmodontia puerulus</i> (Chile; 3, 1)	Muridae	8.0	.18	.018	D	This study (C.Z.)
ET	82	<i>Eligmodontia typus</i> (Argentina; 7, 2)	Muridae	7.8	.18	.016	SD	This study (E.C.-V. and C.Z.)

Table A1 (Continued)

Tip Code	Tip No.	Genus and Species (Common Name)	Family	RMT	Kidney Mass (g)	Body Mass (kg)	Habitat	Reference
Et	83	<i>E. typus</i>	Muridae	11.42		.01858	D	Diaz and Ojeda 1999
em	84	<i>Eligmodontia moreni</i>	Muridae	10.66		.0145	D	Diaz and Ojeda 1999
Eh	85	<i>Eligmodontia hirtipes</i> (Chile; 3, 1)	Muridae	9.0	.18	.017	D	This study (C.Z.)
EM	86	<i>Eligmodontia marica</i> (Argentina; 3, 0)	Muridae	7.5	.24	.017	SD	This study (E.C.-V. and C.Z.)
Sd	87	<i>Salinomys delicatus</i>	Muridae	13.98		.0125	D	Diaz and Ojeda 1999
ar	88	<i>Andalgalomys roigi</i>	Muridae	12.6		.028	D	Diaz and Ojeda 1999
a2	89	<i>A. roigi</i> (Argentina; 4, 1)	Muridae	8.2	.26	.029	SD	This study (E.C.-V. and C.Z.)
AO	90	<i>Andalgalomys olrogi</i>	Muridae	13.48		.023	D	Diaz and Ojeda 1999
gg	91	<i>Graomys griseoflavus</i>	Muridae	9.64		.0545	D	Diaz and Ojeda 1999
GG	92	<i>G. griseoflavus</i> (Argentina; 6, 3)	Muridae	7.1	.966	.056	SD	This study (E.C.-V. and C.Z.)
GD	93	<i>Graomys domorum</i> (Argentina; 3, 1)	Muridae	6.7	.88	.082	D	This study (E.C.-V. and C.Z.)
Ae	94	<i>Andinomys edax</i> (Argentina; 2, 1)	Muridae	6.4	.38	.065	D	This study (E.C.-V. and C.Z.)
IP	95	<i>Loxodontomys pikumche</i> (Chile; 4, 2)	Muridae	5.2	.52	.043	M	This study (C.Z.)
aS	96	<i>Auliscomys sublimis</i> (Chile; 3, 1)	Muridae	7.1	.28	.036	D	This study (C.Z.)
Ab	97	<i>Auliscomys boliviensis</i> (Chile; 6, 3)	Muridae	7.0	.50	.033	D	This study (C.Z.)
PO	98	<i>Phyllotis osilae</i> (Argentina; 3, 2)	Muridae	6.7	.22	.053	SD	This study (E.C.-V. and C.Z.)
pO	99	<i>Phyllotis osgoodi</i> (Chile; 4, 2)	Muridae	7.0	.38	.064	D	This study (C.Z.)
Pv	100	<i>Phyllotis xanthopygus vaccarum</i> (Chile; 2, 4)	Muridae	7.6	1.22	.082	D	This study (C.Z.)
Px	101	<i>Phyllotis xanthopygus</i>	Muridae	10.1		.044	D	Diaz and Ojeda 1999
PX	102	<i>P. xanthopygus</i> (Argentina; 7, 2)	Muridae	7.3	.34	.034	D	This study (E.C.-V. and C.Z.)
pC	103	<i>Phyllotis chilensis</i> (Chile; 3, 1)	Muridae	9.2	.26	.027	D	This study (C.Z.)
pR	104	<i>Phyllotis rupestris</i> (Chile; 5, 2)	Muridae	7.6	.28	.036	D	This study (C.Z.)
pM	105	<i>Phyllotis magister</i> (Chile; 3, 1)	Muridae	7.3	.92	.113	D	This study (C.Z.)
Pd	106	<i>Phyllotis darwini</i> (Chile; 5, 4)	Muridae	6.4	.42	.06	D	This study (C.Z.)
Nm	107	<i>Neotoma mexicana</i>	Muridae	6.3		.098	SD	Brownfield and Wunder 1976
N2	108	<i>Neotoma albigula</i>	Muridae	6.56		.095	SD	MacMillen and Lee 1967
OT	109	<i>Onychomys torridus</i>	Muridae		.3306	.0219	D	Altschuler et al. 1979
P4	110	<i>Peromyscus leucopus</i> (field-caught adult in South Dakota)	Muridae	7.3	.120	.0298	SD	Oswald 1998, and personal communication for body mass
P3	111	<i>P. leucopus</i> (field-caught adult in Pennsylvania)	Muridae	7.28	.117	.0268	M	Oswald 1998, and personal communication for body mass
PL	112	<i>P. leucopus</i> (New York)	Muridae	6.5		.0218	M	Heisinger et al. 1973; Deavers and Hudson 1979
Pe	113	<i>Peromyscus eremicus</i> (cactus mouse)	Muridae		.4454	.0228	D	Altschuler et al. 1979

Table A1 (Continued)

Tip Code	Tip No.	Genus and Species (Common Name)	Family	RMT	Kidney Mass (g)	Body Mass (kg)	Habitat	Reference
Dg	114	<i>Dicrostonyx groenlandicus</i> (collared lemming)	Muridae		.481	.0552	M	Quiring 1950
Na	115	<i>Neofiber alleni</i> (round-tailed muskrat)	Muridae	3.6		.25	A	Pfeiffer 1970; Sperber 1944
OZ	116	<i>Ondatra zibethicus</i> (muskrat)	Muridae		7.45	.9	A	Quiring 1950
Cg	117	<i>Clethrionomys gapperi</i> (southern red-backed vole)	Muridae	4.3		.0248	M	Deavers and Hudson 1979
MA	118	<i>Microtus agrestis</i> (field vole)	Muridae	5.8	.48	.03	M	Sperber 1944
Me	119	<i>Microtus pennsylvanicus</i> (= <i>drummondii</i>)	Muridae	6.1		.038	M	Beuchat 1996
MP	120	<i>Microtus pennsylvanicus pennsylvanicus</i>	Muridae		.31145	.02655	M	Heisinger et al. 1973; Quiring 1950
Mz	121	<i>Mesocricetus auratus</i> (golden hamster)	Muridae		.636	.12	SD	Spector 1956
mA	122	<i>M. auratus</i>	Muridae	8.6	.61	.0593	SD	Brosh 1971, as translated by G. Perry (personal communication)
Ma	123	<i>M. auratus</i>	Muridae	8.01		.105	SD	Munkacsi and Palkovits 1977; Trojan 1977, 1979
Cr	124	<i>Cricetus cricetus</i> (common hamster)	Muridae	5.69	.996	.108	M	Quiring 1950; Sperber 1944; Trojan 1977, 1979
Gs	125	<i>Gerbillurus setzeri</i>	Muridae	8.35		.028	D	Downs and Perrin 1991; Frean et al. 1998
G2	126	<i>Gerbillurus paebe</i>	Muridae	5.97		.029	D	Downs and Perrin 1991; Frean et al. 1998
sc	127	<i>Skeetamys calurus</i> (bushy-tailed jird)	Muridae	9.03	.456	.0577	D	Brosh 1971, as translated by G. Perry (personal communication)
Mh	128	<i>Meriones hurrianae</i> (Indian desert gerbil)	Muridae	12.6		.062	D	Purohit 1975; Goyal et al. 1988
Ms	129	<i>Meriones shawi</i>	Muridae		.627	.1707	D	Rabhi et al. 1996
MT	130	<i>Meriones tristrami</i> (jird)	Muridae	8.2	.378	.0551	D	Brosh 1971, as translated by G. Perry (personal communication); see also Borut and Shkolnik 1974
Mu	131	<i>Meriones unguiculatus</i> (mean of sexes)	Muridae		1.01	.094	D	Kramer 1964
MU	132	<i>M. unguiculatus</i> (Mongolian gerbil)	Muridae	7.51		.07	D	Munkacsi and Palkovits 1977; Edwards et al. 1983
MC	133	<i>Meriones crassus</i> (jird)	Muridae	10.34	.46	.0671	D	Brosh 1971, as translated by G. Perry (personal communication)
Po	134	<i>Psammomys obesus</i> (sand rat)	Muridae	10.7		.15	D	Sperber 1944; Schmidt-Nielsen 1964
Gd	135	<i>Gerbillus dasyurus</i>	Muridae	9.97	.176	.028	D	Brosh 1971, as translated by G. Perry (personal communication)
G3	136	<i>Gerbillus gerbillus</i>	Muridae	10.17	.152	.0215	D	Brosh 1971, as translated by G. Perry (personal communication)

Table A1 (Continued)

Tip Code	Tip No.	Genus and Species (Common Name)	Family	RMT	Kidney Mass (g)	Body Mass (kg)	Habitat	Reference
Gg	137	<i>G. gerbillus</i> (northern pygmy gerbil)	Muridae	10.5	.396	.044	D	Sperber 1944; Burns 1956; Khalil and Tawfic 1963; Schmidt-Nielsen 1964
RN	138	<i>Rattus norvegicus</i> (Chile; 5, 4)	Muridae	6.5	1.78	.17	M	This study (C.Z.)
Rr	139	<i>Rattus rattus</i> (wild-caught rat)	Muridae	5.8		.172	M	Sperber 1944; Collins 1978
Hc	140	<i>Hydromys chrysogaster</i> (water rat)	Muridae	3.9		.85	A	Sperber 1944
Mg	141	<i>Mesembriomys gouldii</i> (shaggy rabbit-rat)	Muridae	8		.65	D	Purohit 1974b
NA	142	<i>Notomys alexis</i> (Australian hopping mouse)	Muridae	7.9	.336	.029	D	Hewitt 1981; MacMillen and Lee 1967, 1969; Purohit 1974a
Ph	143	<i>Pseudomys</i> (= <i>Leggadina</i>) <i>hermannsburgensis</i>	Muridae	7.5		.0126	D	MacMillen and Lee 1967; MacMillen et al. 1972; Purohit 1974a
PD	144	<i>Pseudomys</i> (= <i>Leggadina</i>) <i>delicatula</i>	Muridae	7.7		.012	D	Purohit 1974a, 1974b
aM	145	<i>Apodemus mystacinus</i>	Muridae	6.38	.25	.0364	SD	Brosh 1971, as translated by G. Perry (personal communication); see also Borut and Shkolnik 1974; Shkolnik 1988
M4	146	<i>Mus musculus</i> (Chile; 3, 1)	Muridae	6.6	.08	.012	M	This study (C.Z.)
Mm	147	<i>M. musculus</i> (house mouse, assumed to be wild)	Muridae	8		.018	M	Sperber 1944
M2	148	<i>Mus domesticus</i> (Caithness British Mainland; smallest body mass reported)	Muridae		.1758	.0119	M	Berry and Jakobson 1975
M1	149	<i>M. domesticus</i> (Faray Orkney Island; largest body mass reported)	Muridae		.4791	.0262	M	Berry and Jakobson 1975
Tn	150	<i>Thallamys nigricauda</i>	Muridae	5.67		.082	D	Frean et al. 1998
AN	151	<i>Arvicanthis niloticus</i>	Muridae		1.3	.122	SD	This study (M.A.A.)
Oa	152	<i>Otomys angoniensis</i>	Muridae	4.06	1.21	.1158	M	Pillay et al. 1994, and personal communication
Os	153	<i>Otomys sloggetti robertsi</i>	Muridae	4.3	.96	.1023	M	Pillay et al. 1994, and personal communication
Oi	154	<i>Otomys irroratus</i>	Muridae	3.39	1.76	.1723	M	Pillay et al. 1994, and personal communication
Ou	155	<i>Otomys unisulcatus</i>	Muridae	5.99	1.88	.1176	D	Pillay et al. 1994, and personal communication
pL	156	<i>Parotomys littledalei</i>	Muridae	6.79	3	.107	D	Pillay et al. 1994, and personal communication
Pb	157	<i>Parotomys brantsii</i>	Muridae	5.81	1.75	.112	D	Pillay et al. 1994, and personal communication
PC	158	<i>Praomys</i> (= <i>Mastomys</i>) <i>coucha microdon</i>	Muridae		.265	.0218	M	Quiring 1950
A5	159	<i>Acomys cahirinus</i> (north Israel)	Muridae	8.31	.189	.0366	SD	Brosh 1971, as translated by G. Perry (personal communication)

Table A1 (Continued)

Tip Code	Tip No.	Genus and Species (Common Name)	Family	RMT	Kidney Mass (g)	Body Mass (kg)	Habitat	Reference
A4	160	<i>A. cahirinus</i> (Sinai)	Muridae	9.92	.171	.0329	D	Brosh 1971, as translated by G. Perry (personal communication)
A3	161	<i>A. cahirinus</i> (Kabri)	Muridae	9.32		.0366	SD	Weissenberg 1977
A2	162	<i>A. cahirinus</i> (Eilat)	Muridae	10.84		.0329	D	Weissenberg 1977
Ac	163	<i>A. cahirinus</i>	Muridae	9.4		.033	D	Purohit 1975; Haines and Schmidt-Nielsen 1977
aR	164	<i>Acomys russatus</i> (spiny mouse)	Muridae	11.4	.229	.0528	D	Brosh 1971, as translated by G. Perry (personal communication); see also Borut and Shkolnik 1974

Note. For habitat, A = aquatic, M = mesic, DM = desert mesic (as used by Beuchat [1996]), SD = semidesert, and D = desert. For statistical analyses and figures, species listed as DM were pooled into the SD category. For species new to this study, sample sizes are listed in parentheses (male, female).

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